## Comparison of snow cover April 1 or revised forecast May 1-15 and run-off April-July (per cent of normal)

N. B. Until 1918-19 unrevised snow cover April 1 is used as a forecast. Those revised May 1-15 marked by an R placed before number. Those revised on basis of new data after season was over are followed by Rev., and new estimate in parentheses.

				East s	lope of Sierra				,	West slor	pe of Sierra	
Season	Truckee (exclusi Tahoe), 351,200	Lake Tahoe, feet, 204,180		Carson (but sub heavy divers 251,476 A. F. (N. B.—Courses	ions),	West Walker, 19 A. F. (N. B.—Snow comostly in East W	ourses	South Yuba, 205,4 (Heads against Tr		Mokelumne, 461,48 (Heads against C		
	Forecast		Forecast	Run-off	Forecast	Run-off	Forecast	Run-off	Forecast	Run-off	Forecast	Run-off
1910-11 1911-12 1912-13 1913-14 1914-15	No forecast until 1921-22 except that for adjoin- ing basin of Lake Tahoe.	65. 9 190. 9 52. 2 56. 2 144. 2 92. 7 130. 9 101. 5	82.7 170.4 49.7 58.2 153.8 88.2 { 151.9 {(101.9 Rev.)	172. 3 64. 5 69. 3 150. 6 89. 8	No forecast until 1917-18, but compare adjoin- ing Tahoe for similarity.	64, I 176, 7 42, 4 57, 2 162, 9 93, 3 125, 7 128, 7	No forecast until 1918-19, but no- tice usual simi- larity to Tahoe and Carson, ad- joining basins to north.	96. 6 150. 6 56. 2 50. 9  119. 9 106. 9	(No snow survey until 1915-16. Then only one wind-swept course until 1918-19.  168.4 (148.4 Rev.) 120.7	68. 4 119. 3 68. 3 70. 1 99. 5 109. 8 122. 2 106. 0	Only survey course at Blue Lakes at Crest and interpolation from S. Y u b a. N o te close correspondence between run-off S. Yuba and Mokelumne though the Amercan intervenes.	115.1
1918-19 1919-20 1920-21 1921-22 1922-23	R 96.0	51. 2 73. 7 117. 6 82. 0	\$\begin{cases} 96.2 \\ (56.2 \text{ Rev.}) \\ 80.8 \\ R \text{ 51.3} \\ R \text{ 80.0} \\ R \text{ 121.3} \\ R \text{ 95.1} \end{cases}\$\\ R \text{ -1.9} \\ 80.2	33.6 72.9 56.0 90.4 124.1 94.0 -3.0 101.2	{ 100. 2 (80. 2 Rev.) R 83. 9. R 70. 0. R 103. 0. R 124. 8. R 85. 9. R 26. 0.	66, 6 39, 6 78, 6 121, 2	R 83.0 R 74.8 R 102.0 R 149.3 R 92.0 approxi- mately. R 32.6	70. 0 92. 4 121. 2 85. 3	{ 85. 4 R 89. 2 R 67. 5 R 109. 0 R 141. 8 R 98. 6 R 25. 1 G 29. (75. 7 Rev.)	101. 9 121. 3 1 99. 2	100. 2 estimated (80. 2 Rev.) R 83. 9 R 68. 0 R 103. 6 R 130. 0 R 81. 0 approximately.	81. 6 72. 4 98. 4 126. 9 1 74. 7 24. 1 95. 7

<sup>1</sup> Data for July lacking, making thus only a 3-month run-off. The inclusion of July would decrease the divergence in the case of the Mokelumne.

## AN EXAMINATION BY MEANS OF SCHUSTER'S PERIODOGRAM OF RAINFALL DATA FROM LONG RECORDS IN TYPICAL SECTIONS OF THE WORLD

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[This paper supplements that by the same author in Monthly Weather Review, Oct. 1924]

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By DINSMORE ALTER

[University of Kansas, Lawrence, Kans., Dec. 18, 1925 1]

## SYNOPSIS

This is the ninth of a series of papers on the rainfall of the world, and the second on the application of Schuster's Periodogram. In the last application of this method, published in the Monthly Weather Review of October, 1924, periods longer than nine years were investigated. In this one, periods are examined between nine and two and one-sixth years. In the next paper, which is already mostly computed, still shorter periods will be considered. The aim of these investigations is to examine typical sections systematically, so that all facts concerning rainfall periodicities, which are inherently possible in data at the present time, may be established. It is believed by the author that this question requires such a method as the periodogram, through which periodicities and probabilities are shown, entirely free from the personal bias which must affect the judgment when almost any other method is used. At present, it is his belief, the great need is for such a careful examination of data, rather than for theorizing regarding causes. It is only through thus establishing accurate quantitative relationships that the theories regarding causes can be given the sound footing which they require. Naturally a knowledge of causes is the final goal of all research, but any short cut to theories regarding them is too dangerous to use.

The following summarizes the principal results obtained so far.

(a) Rainfall periods certainly do exist.

(b) There is, in all sections of the world examined, a very marked bias toward harmonics of the sun-spot period, too much so to be merely accidental.

(c) It is impossible to say at present whether these periods are

constant or varying in length, however, the bulk of the evidence favors the former.

(d) It would be too unsafe to make agricultural predictions on the basis of results so far obtained. However, some sections of the world indicate quite strongly that this may be possible in the future.

(e) The more nearly a climate approaches a pure marine the more nearly does its periodogram give us definite results.

# SCHUSTER'S PERIODOGRAM METHOD OF FINDING HIDDEN PERIODICITIES

Schuster's method is the most careful analytical net which has been devised to investigate the existence of periodicities, hidden from casual inspection by means of accidental errors or by the presence of multiple periodicities. Various attempts have been made to use shorter methods of analysis but all these seem unsafe to the writer, some because real periods may be overlooked, others because they permit accidental periodicities to appear real.

Little summary of the method is necessary here, merely a statement of the equations being sufficient. Given data  $q_1 = q_n$ , assume any period  $P_j$  times the datum interval. Let  $\varphi_i$  be the phase angle for the datum  $q_i$ , so that  $\varphi_{i+1} - \varphi_i = \frac{2\pi}{P}$ .  $(\varphi_i = 0)$ 

so that 
$$\varphi_{i+1} - \varphi_i = \frac{2\pi}{P} \cdot (\varphi_1 = 0)$$

Define:

$$A_{j} = \sum_{1}^{n} q_{i} \cos \varphi_{i}; B_{j} = \sum_{1}^{n} q_{i} \sin \varphi_{i}$$

$$I_{j} = \frac{A_{j}^{2} + B_{j}^{2}}{n}; \tan \Phi_{j} = \frac{B_{j}}{A_{j}}$$

where  $\Phi_j$  is the phase of the best sine curve of period  $P_j$  at the instant of observation of  $q_1$ , and  $I_j$  is proportional to the square of the amplitude of this curve. Periods  $P_j$  are chosen of lengths such that there is little phase discourage between edicining area during the state of the square divergence between adjoining ones during the stretch of data, and  $I_j$  is computed for each. A curve is then drawn with P's as abscissæ and I's as ordinates.

 $<sup>^{\</sup>rm I}$  Since sending the manuscript for publication, an excellent article by Sir Gilbert T. Walker on the periodogram has appeared in No. 216 of the Quarterly Journal of the Royal Meteorological Society. Our conclusions regarding the strength and limitations of the method parallel each other very closely although in general his treatment is the more elegant.—D. A.

Usually a quantity  $H_j = I_j/I_{mean}$  is plotted instead of  $I_j$ . Schuster determined the mean I by measurement of the area under the curve. In one of my papers (1f) I have computed this mean value by the equation

$$I_m = 1.099\epsilon^2$$

where  $\epsilon$  is the probable error of one datum under the assumption that all their deviations are accidental. If the order of the deviations of the q's is not accidental, we have peaks in the periodogram higher than would be expected from the theory of probabilities. The frequency of distribution of such peaks, under the law of accidental grouping, is e<sup>-n</sup>. It is obvious that peaks higher than would be expected from error distribution, will raise the mean height of I, if we have only a limited number of values of q, or a short stretch of the periodogram. The computed value of  $I_m$  is that which we would have obtained by measurement, if we had had many data and had used a long stretch of periodogram. It is, therefore, not only more convenient but also more accurate to use in computing probabilities. After publication of the above equation, I found a more convenient form for computation:

$$I_m = \frac{\sum_{i=1}^{n} \sigma_i^2}{2(n-1)}$$

where  $\sigma_i$  is the deviation of  $q_i$  from the mean q. As we use smaller and smaller values of P, we find that the amplitude of a computed period is less than it would have been had we used shorter datum intervals and, therefore, a larger P for the same period. However, it is easy to reduce a computed I to what we would have obtained from the shorter intervals, by means of the equation (1f)

$$I^{1} = I \frac{(x-y)^{2}}{4 \sin^{2} \frac{1}{2}(x-y)}$$

where x is the phase of  $q_i$  and y that of  $q_{i+1}$ , expressed in radians. It is almost needless to remark that computations of probability must be made from an H which has not been multiplied by this factor. Nevertheless, the factor has some real value, since it gives us the most probable values of the intensities and amplitudes of the best sine curves of periods  $P_i$  and enables us, thus, to compare their effect on the data. This factor is given as column F of Table 5. Probably it would have been better to have plotted the periodograms after multiplication by F; however, I have used the original values for this in order that the graphs may conform to long established custom.

With periods of three times the datum interval or less, F begins to get large and accordingly the ratio of accidental error to intensity of any real period of a given amplitude becomes greater. For these reasons any such real periodicity will be displaced more from its true value, both in length and intensity, than will longer periods of the same amplitude. It is, therefore, impossible to demand as good an agreement as would be expected of longer periodicities. In this paper periodicities have been investigated, using yearly datum intervals, for the whole range between nine and two and a sixth years. However in the next paper, now more than half finished, half yearly datum intervals will be used. That paper will cover the range two and a half to one and

a twelfth years, thus overlapping the most inaccurate part of the present periodograms.

Limitations and powers of the method.—There are a few of these which it may be well to mention, although probably almost all of them are well known to everyone who has studied the method. Most of them have been discussed in detail in various publications.

1. No matter how small amplitudes of real periods may be, they can be shown definitely to be real, if we have enough data. In oral discussion of a paper read recently, the objection that this is not true was raised against the method. Schuster's method yields nothing to any other method in showing small real periodicities. The objection has arisen through the fact that some other methods do not show clearly enough the lack of evidence in favor of these periods.

2 Periods of large amplitude will have both their length and intensity most accurately shown, since the greater they are, the less is the ratio of accidental errors to them. For this reason if we have any grounds, either theoretical or statistical, to suspect a given set of periods, as for instance harmonics of the sun-spot period, we will demand of the highest peaks a very much greater coincidence with these harmonics than for lower peaks, even

though these also be high enough to indicate a good

possibility of reality.

3. If two stretches of data are investigated, the longer including and being a continuation of the shorter, intensities of periods should remain the same, on the average for the two stretches, if they be accidental but should be larger for the longer stretch, if they be real. This refers, of course, to periodograms plotted from the ratio H<sub>i</sub>.

4. In determining reality of periods, not only the intensity, but also the length of the period with respect to some other plausibly related phenomenon should be considered. For example, if in this rainfall investigation, peaks of medium heights, nearly at harmonics of the sun-spot period were to be found, it would be legitimate to regard them as more probably real than we would regard those of the same intensities but whose lengths of periods had no special significance. However, it is mainly a matter of judgment and personal opinion what weight shall be attached to this consideration, unlike the matter of intensity for which there is an accurate mathematical probability. For this reason extreme caution must be used with this argument. It can be used merely as an additional evidence to the primary one derived from the intensity.

5. The same period, found in independent records of any one kind of data, is almost as strong an argument for reality as is intensity. This is especially true for chronologically different records.

6. In the preceding paper (1d) it was shown that the accuracy with which any period is located is less than that which would be expected from casual examination

of the periodogram.

7. The expression "expectancy ratio under error law" would be more accurate to use than "probability," since the calculation of  $e^{-\pi}$  shows the ratio, to be expected by accident, of peaks of a given height to the number of peaks computed. The probability, based on mere statistics, that the peak represents a real period is much less than this ratio, Also each peak, established definitely, makes minor peaks of medium height more worthy of consideration. Although Shuster was very emphatic in stressing this point about the probabilities of reality, it seems not to have been appreciated by some.

· He considers that only those peaks for which this expectancy ratio is less than one in two hundred are worthy of consideration as possibly real, when based on statistics This judgment seems quite sound, although the number of points computed in the periodogram should be considered, and it will be adopted here as a criterion, except when modified by 4 and 5 above. In such cases a larger ratio may be considered as sufficient to indicate a possible real periodicity.

8. In using Schuster's periodogram there is absolutely no danger of prejudicing the solution in favor of some particular value, as has been done by other methods.

9. The method can be adapted to investigation of variable periodicities. The same limitations apply here as to other methods, although, as in 1 above, they are more obvious than in most other methods which have been applied to such cycles. In order to make a legitimate examination, the law of variation must be assumed from hypotheses other than an examination of the data. For example, it is entirely legitimate to crowd up or stretch out weather data in accordance with the apparent variations of the sun-spot period before applying analysis to them. But it would be entirely forbidden to take these equal phase intervals for the sun-spot data themselves. Also, similarly, it would be improper to look at weather fluctuations and say that when crests were far apart a period had lengthened, merely from an examination of these data themselves.

10. In computing the mean height of the curve the tota data are used. Since, in order to hold approximately to complete cycles, some data are usually discarded at one end of the stretch, it would be most strictly correct to compute  $I_m$  for each point, only from the data used for that point. This would involve considerable labor for a very slight improvement in the value of H. The neglect will always lower H slightly and, therefore, merely results in probabilities of reality being actually a little greater than we have computed. Schuster discusses this near the bottom of page 74 of the reference (2c) above.

11. For short periods, such as investigated in this paper, it is no longer permissible to abbreviate the work by repeating or averaging a month every now and then to get fractions of years, as was done in the last paper. work is enormously increased by the fact that the mean phase of any year must be accurately computed and that only in comparatively few cases will any phase angle be repeated more than twice during the stretch. Instead of a sum or a mean being multiplied by sine, and cosine, each value of  $q_i$  must be so multiplied. However, if we assign the same number of years to the stretches of data from different parts of the world, the phase angles, sines, and cosigns, once determined, may be used for all. In this case there were 73 years of data for the Pacific coast of the United States, and this number was used for all other sections, except the Punjab where only 62 years are available.

# EXAMINATION OF DATA FOR VARIABLE PERIOD

An hypothesis which has been discussed somewhat in recent years is that weather periods or cycles do exist and that they stretch out or close up so as to keep in step with the variations of the sun-spot period. For years this period was considered to be 111/8 years, which is the mean visible period between successive maxima or minima of the number of sun spots. Recent work by Hale (3) at the Mount Wilson Observatory shows that the period of variation of magnetic polarity is exactly double this

and that it is better to consider the mean period as being 22.25 years.

In the earliest papers of this series, a short period was examined on the hypothesis of a forced phase agreement between rainfall and sun spots. There the datum intervals were months and it was difficult to make the proper table for expanding or contracting the number of data to keep a constant number of phase steps between successive sun-spot maxima or minima. Here, using yearly data. the problem is much simpler. Table 4 shows the years to be repeated or averaged to force such a relationship. Within narrow limits, this choice of years is arbitrary. However, a rule of spacing as uniformly as possible leaves little choice. The method is, of course, but an approximation; nevertheless, if the weather cycles exist and do thus change their period, the periodogram derived from the data thus adjusted should show higher peaks than from the unadjusted. This adjustment gives exactly 22 datum intervals to the sun-spot period, instead of the average of 2214 as for the unadjusted. The writer was surprised to find how little difference there is in the tables of adjusted and unadjusted data since 1850, the year for which the 73 data for these investigations usually begin. Twenty of the years agree exactly in both the unadjusted and adjusted tables, 47 differ by but one place, and only 6 by as much as two places. It is evident that there will, in general, be a great similarity between the periodograms and that it will be very difficult to tell whether periods approximately constant in length or changing with the apparent sun-spot variation are the more probable. the earlier sun-spot data, which show large deviations from the mean period, can be accepted as approximately accurate, then the preceding 73 years, for which we have data from Northern Europe, should tell us much about this question.

## THE DATA USED IN THIS PAPER

(a) Pacific coast of the United States.—These are identical with Table 5 of the preceding paper and, therefore, will not be reproduced here.

(b) Northern Europe.—Many new data have been added so that they are given in toto as Table 1.
(b) The Punjab of India.—These data are identical with Table 6 of the former paper, except for the addition of the years 1919-1924. Table 2 shows these later years only.

(d) Eastern United States.—In the main, these are the same as Table 4 of that paper. New England stations have been added. Table 3 shows only these additional stations and the means of these with the stations of the

The Pacific coast of the United States.—The results for this section are shown in the first columns of Table 5 and in Figure 1. Three periods stand out above all others in the unadjusted periodogram. First is one of H=8.98 at P = 2.469 years; second H = 7.42 at 5.38 years, and third is H=7.17 at 4.42. The computed expectancy ratios for these peaks follow. For the largest value of H, one out of every 7,950 should be of this height by mere accident. In the periodogram there are 86 computed points, with two of this height. It would be, therefore, entirely improbable that we would obtain this peak by accident. For the next two peaks the ratios are 1 to 1,660 and 1 to 1,280. If the sequence of deviations on the Pacific coast is but accidental, one would be much surprised to obtain any peaks as high as this, and much more surprised to find three. One-ninth the sun-spot period is 2.472 years,

an agreement with the highest peak more perfect than one could possibly expect, indeed far within the accuracy with which the peak can be located. One-quarter of the sun-spot period is 5.56 years, differing from the second of the observed peaks by 0.18 of a year. The computed uncertainty in the position of the peak is much larger than for the shorter period, having increased both because of the lesser number of cycles in the 73 years' data and also because of the lesser phase change in one year. It is 0.14 of a year, approximately equal to the discrepancy. Moreover, the steepness of the two sides of the peak indicates that, if more points had been computed, the crest would have fallen to the right of its present position, somewhere between 5.40 and 5.45, giving a smaller discrepancy. One-fifth the sun-spot period is 4.45 years, 1.11 years less than the fourth harmonic. Therefore, the peak agrees with the harmonic to better than one-sixth the interval between harmonics, an agreement closer than would have been expected by accident, but not impossibly accidental. The third peak is at 4.42 years, which differs from its harmonic by the almost negligible quantity of 0.03 of a year. Therefore, in this periodogram we have three peaks so high that we would expect none of them by accident, two of them in almost perfect coincidence with sun-spot harmonics and the third closer than we would expect through chance.

Two other peaks are found just at the limit of the Schuster criterion, and because of the presence of the very high ones, they become worthy of some notice. The higher is at 2.25 years, differing from the tenth harmonic by 0.02 of a year. The next is at 3.17 years, as perfect an agreement with the seventh harmonic as the solution permits. The lowest of the highest six peaks is at 6.83 years and is the first peak to diverge seriously from the sun-spot harmonics. Each of the highest five peaks fall remarkably close to the harmonics of the sun-spot period.

Examination of the adjusted periodogram shows the expectancy ratio of the highest peak to be one in 5,500. The peaks, although still high, average lower and the coincidence with sun-spot harmonics is lacking. Therefore, so far as we can tell from the available data of this section, constant periodicities, at least so far as length of period is concerned, are the more probable. Some of sort periodicity almost undoubtedly exists and there is a quite probable relationship to the sun-spot period. This section has the purest marine climate of any of those investigated.

Northern Europe and the British Isles.—In this section, which is next nearest to being a pure marine climate, 146 years of data have been used. The first pair of periodograms have been computed from the years 1777–1849 and the later from 1850–1922.

The 1777–1849 unadjusted periodogram shows but two peaks of much interest, however, one of these is far the highest peak found for any section. For it, H=16.95 and it is found at 2.449 years, almost exactly where the highest peak was found for later years on the Pacific coast. The expectancy ratio of peaks of this height is one in 22,400,000. Independently of the fact that it is at one of the sun-spot harmonics and of the fact that it agrees almost perfectly with the highest peak of a different section and from a later stretch of years, there is little question that this peak is not accidental. A period equal to one-ninth the 22.25 year sun-spot period actually did exist in northern European rainfall during these years.

The second highest peak has an expectancy ratio of 1 in 1,600. It falls at 4.17 years, which is 0.28 of a

year less than the fifth harmonic, which was found in the data of the Pacific coast. The third peak is much lower with an expectancy ratio of one in 250. Its position resembles somewhat that of the seventh harmonic, also found in Pacific Coast data, but there is little to depend on, either from its magnitude or position. Of course, if it actually be a real peak, it will, due to its small magnitude, be subject to greater displacement than higher ones.

When we turn to the adjusted data we find once more that the peaks are lower, although much higher than accident would place them. The expectancy ratio of the highest is one in 4,370, of the second highest one in 1,280 and of the third highest is one in 200. Again we find that the adjusted peaks bear no relationship to the sun-spot period. This is extremely important evidence in favor of nonvarying periodicities, for it was during this epoch that the sun-spot period appeared to diverge most from constancy. Although, for the latter stretches of data we would expect the false hypothesis to show nearly as well as the true, here as we would expect, we do find the differences of the two periodograms to be very marked.

So far all our evidence has been extremely favorable to an hypothesis of constant periodicities (at least so far as the length of the period is concerned) which occupy certain harmonics of the sun-spot period. However, the data from Northern Europe for the years 1850–1922 tell a different story. The unadjusted shows, it is true, three peaks higher than we would expect from accident, but they are low compared to those of the preceding periodograms. The expectancy ratios are one in 420, in 340 and in 220. The two highest of these fall very nearly at sunspot harmonics, the higher missing the fourth harmonic, also found in the Pacific coast, by only 0.06 of a year, which is practically perfect agreement for periods of this length, and the next missing the seventh harmonic by 0.05 of a year.

When we turn to the adjusted period we find one peak with an expectancy ratio of one in 1870, and two others of about one in 200 each. Of these three only one, and that one of the two lowest, falls near a sun-spot harmonic. That one is very close to the sixth.

This reversal of previous results is surprising. However, an analysis of Table 1 gives us some indication of what has happened. In the data of the later years, a number of new stations have been added, as they began to make records, in an attempt to eliminate, as far as possible, accidental errors and the effects of local storms. Several of these were in Germany, two of them being possibly too far south from the north coast to be true marine rainfall. It is a natural result of the prevailing westerlies, that we can go farther inland for marine type stations on west coasts than on others, especially the east and north. The principal effect of these inland stations comes in the later data, so that, if there be a phase difference between marine and continental stations, these records would cut down peaks instead of reinforcing them.

If we will choose carefully as pure a marine type of climate as is possible in this section, the ninth harmonic, which has disappeared, should reappear if this be the true explanation. A periodogram was computed, therefore, for the years 1850–1922 from the data of the British Isles. Possibly it would have been better to include the records of western France, of Sweden, of Denmark, of the Netherlands, etc., which had already been used in the early curve. If I ever repeat this section I shall do this, especially for a computation of a short periodogram

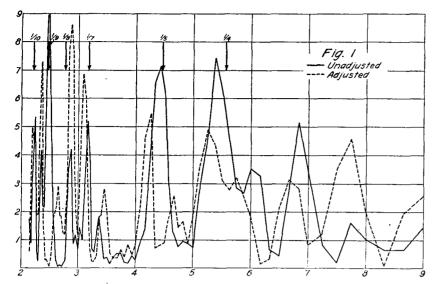


Fig. 1.—Rainfall periodogram, Pacific coast of United States

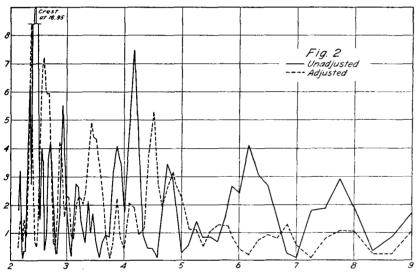


Fig 2.—Rainfall periodogram, northern Europe, 1777-1849

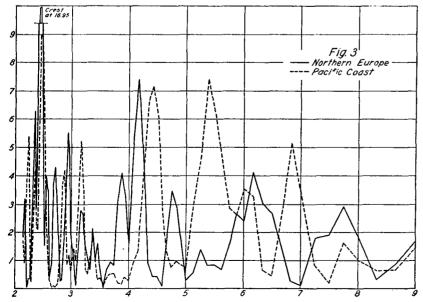


Fig. 3.—Rainfall periodogram, northern Europe, 1777–1849, and Pacific coast of United States, 1850–1922

in the neighborhood of the ninth harmonic. This special periodogram for the British Isles was carried through only for the unadjusted data. A considerable improvement was found. Peaks exist of H=6.97 at P=2.84 years, of 6.81 at 3.17 years, with a secondary of the latter of H = 6.24 at 3.38 years, and of 6.30 at P = 4.25years. There are two minor peaks H=5.18and 4.54, respectively, the latter at 2.42 years, 0.05 of a year from the true position of the missing ninth harmonic. These three highest peaks are all higher than any in the previous unadjusted periodogram and are surpassed by but one peak of the adjusted periodogram. The expectancy ratios are one in 1,300, in 900, and in 600. They hold rather closely to the eighth, seventh, and fifth harmonics, especially to the seventh, for which the agreement is perfect.

It is evident that the exclusion of the inland data has made a considerable improvement, but the closeness of the agreement between the periodograms of 1777–1849 for Northern Europe and of 1850–1922 for the British Isles, which are a large part of the former, can show best only by an examination of the superimposed curves. Quite apparently the main differences are in magnitude only, and we have a very similar "spectrum" from the two epochs. This point will be discussed later. These superimposed curves are shown as Figure 7. Figure 6 shows the two unadjusted Northern Europe periodograms and Figure 3 compares the early Northern Europe with the Pacific Coast.

Northern Europe with the Pacific Coast.

The Punjab of India.—We have one section which is almost as pure a continental type as is to be found. This section is The Punjab, a thousand miles inland and with light winter and heavy summer monsoons. Unfortunately there are only 62 years of data available. This fact is certain to give us smaller values of H, if the peaks be real. If accidental, their mean heights should be unaffected. Tentatively we shall study peaks lower than we demanded for the other sections.

the other sections. In the unadjusted periodogram we find for the highest peak H=4.78 at P=2.78, which is exactly the eighth harmonic. The expectancy ratio of this peak is one in 120. For the second highest peak H-4.49 at P=7.5, with the steepness indicating the true crest between the computed points of  $7\frac{1}{4}$  and  $7\frac{1}{2}$ . This is exactly at the third harmonic. The third highest peak is H=3.60 for P=3.17, almost exactly at the seventh harmonic. Although not as strong evidence of reality as for other sections, because of the low heights of these peaks, this series of agreements is among the prettiest things seen in the investigation.

In this section we find that the adjusted peaks are somewhat higher than the others, with  $H=5.38,\ 5.35,\ 5.07$  and 4.81. The expectancy ratio of the highest peak is one in 215. This peak does not match at all with the harmonics. The second highest at P=2.25 matches the tenth fairly well. The third peak at P=3.75 is very close to one-sixth of 22. The fourth

peak is at 7½ years, very near the third harmonic. On the whole we find, for this section, that the evidence is slightly in favor of the variable period, although not nearly so strongly as is the reverse in the case of the Pacific Coast and Northern Europe.

Eastern United States.—If this section were to be computed again, I would choose only a small part of it, probably New England. The same error was made as in the case of Northern Europe. Data are included from a large region, extending from New England to St. Paul, then to New Orleans and east to Florida. On the whole, it tends to be continental in rainfall.

The highest peak on the unadjusted curve is a symmetrically shaped one, H=5.73 at P=7.5, the third harmonic. Its expectancy ratio is one in 310. The next highest peak is H=5.63 at P=4.75. This is one of two adjoining peaks, the other being H=4.40 at P=4.33. The curve does not get down to normal between them. Neither is at a harmonic, for they straddle P=4.45, the fifth harmonic. The only other point worthy of mention is H=4.50 at P=3.17, the seventh harmonic, which has been so persistent in various parts of the world.

The adjusted curve gives us but one peak, a high one, H=7.78 at P=7.25, the third harmonic of 22. This one high peak, with expectancy ratio one in 2,400, makes this periodogram very striking. However, when all is balanced it seems that the evidence from this section scarcely favors one hypothesis more than the other. Probably it is slightly in favor of the variable period.

# THE BIAS OF THE UNADJUSTED DATA TOWARD SUN-SPOT HARMONICS

We have constantly seen the agreement of peaks of the unadjusted periodograms with harmonics of the sun-spot period. In each section, without a single exception, the highest peak is almost exactly at one of the sun-spot harmonics. This bias continues, in general, to the second and even to the third highest peaks. The following tabulation exhibits clearly how unusual this coincidence actually is.

Section	Highest, H	Peak, P	Нагиопіс	Derived sun-spot period
Pacific Coast	8. 98 16. 95 6. 04 6. 97 4. 78 5. 73	2. 47 2. 45 5. 62 2. 84 2. 78 7. 50	9 9 4 8 8	Years 22, 23 22, 05 22, 48 22, 72 22, 24 22, 50
Mean				22. 37

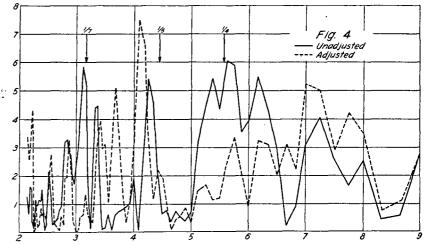


Fig. 4.—Rainfall periodogram, northern Europe, 1850-1922

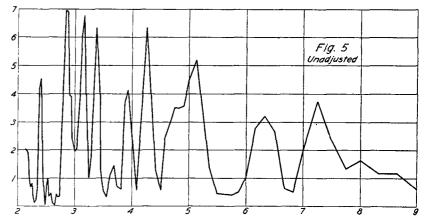


Fig. 5.—Rainfall periodogram, British Isles, 1850-1922

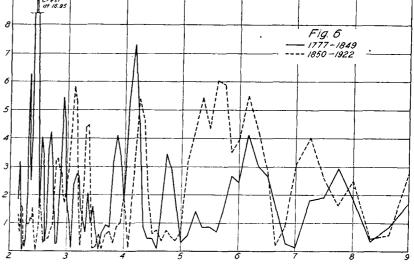


Fig. 6.—Rainfall periodogram, northern Europe, 1777-1849 and 1850-1922

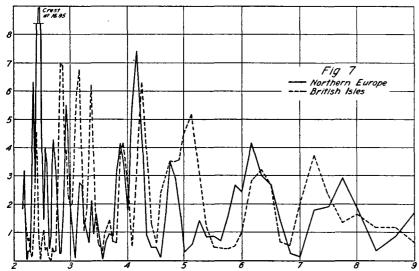


Fig. 7.—Rainfall periodogram, northern Europe, 1777-1849, and British Isles, 1850-1922

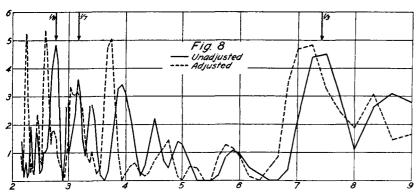


Fig. 8.-Rainfall periodogram of the Punjah

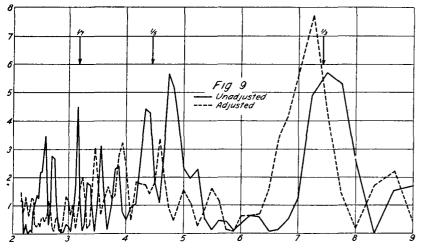


Fig. 9.—Rainfall periodogram, eastern United States

The sun-spot period as computed from the rainfall data disagrees by only 0.12 years from the value obtained from the spots themselves. Its probable error is 0.07 of a year. Regardless of the interpretation which one may put on the periods found here, it seems impossible that the relationships between sun-spot and rainfall periods can be accidental.

#### RECAPITULATION AND CONCLUSIONS

(a) The higher peaks found in the periodograms can not be due merely to accident.

(b) On account of the little difference between the supposedly variable sun-spot period and a constant one, during the last three-quarters of a century, it is impossible to determine definitely whether the periods are fixed or variable, but the bulk of the evidence favors the fixed periods.

(c) The periods are, for some reason, closely related to the sun-spot period. This paper is statistical and does not enter into causes.

(d) The effects seem most pronounced for marine climate and especially so for the pure marine climate of our Pacific coast. This is exactly the result found several years ago in an investigation of a short period (1c).

an investigation of a short period (1c).

(e) Periods of practically constant length, but possibly with varying amplitude, seem most probable. For an identical conclusion regarding sun spots, by Schuster, see pages 89-95 of (2c) in the bibliography.

(f) Nothing has yet been found of sufficient accuracy to use as a basis for long range agricultural forecasts, although the results distinctly encourage the hope that this may be found in the future, at least for the Pacific coast of the United States and perhaps for the Punjab.

(g) For the same reasons that these periods gave very much more definite results than the longer ones of the previous periodogram investigation, it can be expected that the next paper on still shorter periods will be even more definite.

(h) There has been for many years much theorizing regarding causes of supposed relationships. Although the end of all research is to find causes, it seems to the writer that our present need is to establish statistically and accurately the quantitative relationships between solar and terrestial phenomena, in order that there may be a firm basis for the hypotheses of the future.

I wish to thank Professor Marvin, Chief of the United States Weather Bureau, and Professor Talman, librarian, for giving me full access to all stacks and records during three weeks spent at the bureau a year ago. Part of the computations for this paper have been made through a grant from the research committee of the Graduate School of the University of Kansas. Also I am much indebted to Professor Kester for his continued interest in the problem and the sound advice which he has often given. He has carefully studied even details of each paper published by me on this subject during the past five years and many points have been improved and added through his suggestions.

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Table 1.—Northern European Rainfall—Yearly percentages of normal at available stations with long rainfall records

	Compilation of data for 8 English stations	Paris	Berlin	Mean		Compilation of data for 8 English stations	Paris	Berlin	Mean		Compilation of data for 8 English stations	Paris	Berlin	Mean		Compilation of data for 8 English stations	Paris	Berlin	Mean		Compilation of data for 8 English stations	Paris	Berlin	Mean
Normal		496. 9 mm.	549 mm.		Normal		496, 9 mm.	549 mm.		Normal		496. 9 mm.	549 mm		Normal		496. 9 mm.	549 mm.		Normal		496. 9 mm.		
Year 1689 90 91 92 93 94 95 96 97 98		103 128 78 123 124 64 107 106		103 128 78 123 124 64 107 108	Year 1700 1 2 3 4 5 6 7 8 9		109 117 89 94 108 75 83 98 100 119		109 117 89 94 108 75 83 98 100 119	Year 1710 11 12 13 14 15 16 17 18		86 137 115 112 81 95 78 96 72 56		86 137 115 112 81 95 78 96 72 56	Year 1720 21 22 23 24 25 26 27 28 29	109 104 114 103	93 69 79 46 67 93 62 74 88 93	87 77	93 69 79 46 67 93 86 89 96	Year 1730 31 32 33 34 35 36 37 38	86 66 96 84 118 102 111 109 73	87 55 89 42 95 75 82 86 80	121 79 90 109 100 112 77 115 81	98 63 92 78 104 96 90 103 78

Table 1.—Northern European Rainfall—Yearly percentages of normal at available stations with long rainfall records—Continued

Table 1.—Northern European Rainfall—Yearly percentages of normal at available stations with long rainfall records—Continued

	Compilation of data for 8 Eng- lish stations	Edinburgh	Kendal	Greenwich	Chilgrave	London	Haverford West	Glengyle	Belfast	Punq	Abo	Warsaw	4 stations in Norway	Сореправеп	Utrecht	Montdidier	Paris	Lille	Brussela	Koenigsberg	Tilsit	Berlin	Danzig	Mean
Normal.		25.9 in.	52.1 in.	24.7 in.	34.3 in.	25. 6 in.	48. 0 in.	91. 8 in.	34. 6 in.	17. 2 in.	592.5 mm.	575 mm.	49. 7 in.	22. 7 in.	28. 5 in.	(?)	496. 9 mm.	691 mm.	742 mm.	659 mm.	661 mm.	549 mm.	546 mm.	
91 92 93 94 95 96 96 97 98 99 1900 1 1 2 3 4 5 6 6 7 8 8 9 1910 11 11 12 13 14 15 16 16 17 17 18 18 18 18 18 18 18 18 18 18 18 18 18	108 98 106 99 99 99 99 118 87 90 107 89 118 85 97 108 90 130 98 91 138 101 74 88 93 97 80 102 102 102 103 90 103 90 104 105 106 107 108 108 108 108 108 108 108 108	90	104 111 89 103 88 108 95 113 84 112 93	107 746 81 103 113 880 125 76 1313 880 125 76 135 91 994 902 711 122 96 79 94 881 121 77 96 881 124 108 889 75 90 121 127 120 104 102 893 106 80 116 124 108 89 75 90 121 104 102 89 101 104 102 89 104 114 107 104 109 80 114 109 80 114 109 80 101 90 80 114 90 80 101 90 80 114 90 90 90 90 90 90 90 90 90 90 90 90 90	96 96 96 97 98 118 82 90 101 76 131 86 91 131 64 84 84 84 89 91 127 80 91 131 164 81 81 111 176 187 80 91 187 80 91 187 80 91 187 80 91 187 80 91 187 80 91 187 80 91 187 80 91 91 91 91 91 91 91 91 91 91	126 87 74 111 102 116 95 57 109 98 81 110 116 116 88 99 118 88 91 119 119 119 119 119 119	78 74 74 81 104 85 106 88 88 104 95 84 118 92 93 93 105 92 93 111 105 103 91 111 105 109 89 89 104	103 122 114 115 188 779 1104 1107 117 117 117 118 118 1105 1116 1113 1116 1117 1117 1117 1117 1117	109 101 106 109 104 110 100 124 97 94	77 91 140 (75) 95 77 113 118 81 73 118 81 73 118 82 121 146 75 76 81 13 121 121 121 121 121 121 121	92 102 96 103 85 97 98 81 117 102 84 89 103 97 73 104 86 103 97 72 102 103 97 72 103 97 72 103 98 86 103 86 103 87 73 104 86 86 105 86 86 86 86 86 86 86 86 86 86 86 86 86		108 111 104 100 106 104 106 108 86 106 81		92 110 98 128 100 99 97 102 97 102 97 110 92 132 78 110 94 89 95 103 116 89 123 123 121 121 121 121 121 121 121 121	99	98 87 92	104		94 104 111 122 90 107 98 68 84 772 86 68 84 772 86 105 73 91 125 91 65 96 83 84 81 110 94 124 122 110 111 117 108 118 119 119 1111 1117 103 1118 1119 1119 1119 1119 1119 1119 111	81 95 70 70 70 106 94 94 95 102 102 103 109 105 102 102 104 87 66 69 90 103 103 103 103 103 104 105 102 104 105 102 104 105 102 104 105 106 107 108 109 108 109 109 109 109 109 109 109 109	89 110 104 79 91 111 103	1111 777 107 107 108 107 109 101 119 1101 1115 1115 1115 1115 11	99 88 108 89 92 113 95 106 93 109 106 98 102 98 102 98 101 101 101 100 100 100 100 107 96 102 107 97 108 104 98 102 97 107 97 108 101 107 97 108 109 100 98 107 97 102 91 103 104 97 107 97 108 109 100 98 107 97 102 91 103 104 97 107 97 108 109 100 98 107 97 108 109 100 98 100 99 112 100 99 99 112 100 99 99 112

TABLE 2.—The Punjab

[Table supplementary to that published in M. W. R. Oct. 1924, p. 485]

Year	Per cent of normal	Year	Per cent of normal
1918	47 83 52 72	1922 1923. 1924.	71 91 90

Table 3.—Eastern United States

[The mean includes stations of Table 4 of Mo. Wea. Rev., October, 1924, p. 485]

Year	Boston	Lowell	New Bedford	Provi- dence	Mean
1817			94		108
1818	98		88	·	89
1819	81		86		80
1820	101		87		102
1821	84		99		98
1822	62		90		74
1823	107		130		113
1824	82		102		95
1825	81		82		71
1826	94	78	119		89
1827	112	125	136		113
1828	74	91	85		87
1829	107	89	142		107
1830	98	103	140 133		102 104
1831 1832	118 107	128	107	89	104
1833	87	106	92	77	96
1834	91	177	96	95	85
1835	87	78	102	70	90
1836	93	86	93	86	100
1837	77	74	85	72	89
1838	97	91	83	86	95
1839	94	92	96	83	92
1840	112	93	107	93	97
1841	108	97	110	108	105
1842	89	93	85	85	98
1843	107	95	110	96	105
1844	86	86	88	79	88
1845	106	94	104	98	94
1846	69	68	75	69	101
1847	107	112	99	110	109
1848	94	102	88	92	96
1849	92	101	79	79	96
1850	123	123	136	116	115
1851	101	110	112	98	89
1852	110	103	100 83	87 121	100
1853 1854	112 104	106 102	116	105	96 95
	104	108	89	88	98
1855 1856	119	102	80	93	88
1857	116	119	94	101	105
1858	120	86	95	101	104
1859	130	115	111	102	113
1860	118	113	86	87	93
1861	114	104	100	100	97
1862	140	107	94	114	104
1863	155	126	98	125	106
1864	113	92	89	83	91
1865	109	90	100	101	112
1866	116	92	87	104	98
1867	127	110	102	107	110
1868	147	116	122	121	115
1869	151	114	108	110	109
1870	137	112	102	111	102
1871	103	107	107	108	101
1872	115	107	103 112	110 119	100
1873	125	96 86	107	98	115 103
1874 1875	97 115	96	107	118	106
	112	109	91	114	112
1876	118	99	102	110	106
1878	150	137	109	119	117
1879	102	105	92	92	96
44.4			·		30

TABLE 3.—Eastern United States—Continued

[The mean includes stations of Table 4 of Mo. Wea. Rev., October, 1924, p. 485]

Year	Boston	Lowell	New Bedford	Provi- dence	Mean
1880	85	85	87	94	100
1881	120	104	85	101	102
1882	100	99	90	102	103
1883	81	96	94	90	108
884	112	113	119	110	107
1885	103	117	80	90	101
886	96	111	108	118	102
887	77	126	112	115	98
1888	105	143	119	144	111
[889]	91	100	114	127	107
890	- 89	118	134	115	110
891	91	82	104	120	99
892	85	103	93	85	97
[893	96	104	109	116	103
1894	84	81	99	96	9:
l895	92	92	90	115	88
.896	86	100	103	104	94
.897	93	100	110	108	103
898	114	130	136	144	109
899	79	88	96	112	9:
900	101	126	96	108	9-
901	111	130	112	118	91
902	78	124	98	109	100
903	96	100	103	107	103
904	91	96	108	107	9:
905	73	90	89	94	98
906	93	101	93	109	100
907	86	93	99	108	97
908	69	75	91	96	86
909	93	84	92	76	94
910	65	69	82	78	88
911	82	83	91	83	99
912	79	86	99	87	100
913	87	86	99	84	94
914	78	67	84	67	l 8€
915	89	92	95	77	102
916	85	97	100	78	92
917	89	77	84	82	86
918	79	85	71	85	88
919	98	86	102	100	105
920	105	107	108	101	103 94
921	98	102	80	81	
922	94	122	80	102	99 92
923	77	104	68	92	
924	80	91	80	76	94
Normals	43.75	41. 49	46, 21	44. 16	

Table 4.—Years to be repeated or averaged to form variable table, in forced step with sun-spot numbers

To be repeated—										
1751 1754 1762 1765 1767	1772 1776 1780 1807	1809 1830 1831 1832	1840 1868 1885 1920							
,	ro be av	eraged—								
1756- 1759- 1789- 1792- 1795- 1799- 1801- 1804- 1811-	-60 -90 -93 -96 1800 -02 -05	1814 1824 1827 1844 1850 1874 1880 1891 1902	25 28 45 51 75  -81 92							

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	States	sted	-	<u>48288</u>   <u>9859</u>	728428 78458 78646	-24 -24 -20 -20 -20 -20 -20 -20 -20 -20 -20 -20	24 0.1.1.0 24 0.1.1.0	1132448 11111111111111111111111111111111111	78888 78888 78888	52888 440944	252 253 253 212 212 213 213 213 213 213 213 213 21	28885 9855 06490	85 22 10 1 2 2 2 1 0	34.1. 34.1.	27.00.00.00.00.00.00.00.00.00.00.00.00.00
	g S	Adjusted		25.0.4 6.0.1.2.0	2317. 1274. 1023.	210.0.1 190.0.0 20.0.0	24.84 7.0.0 7.0.0 7.0.0		_ c1 cc <u></u>		84444	52 3. 1001.	245 245 245 245 245 245 245 245 245 245	2883 1.0.1.1	
	United	#	<u>                                     </u>	82822	35 11 25	350313	8732	\$6248	82788		8 2 2 5 8 3 8 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	00 20 20 20 20 20 20 20 20 20 20 20 20 2	223002	652334	22222
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		Unadjusted			163 5. 7 140 4. 9 36 1. 2 150. 5	30000 00000 13800	130.4 130.4 150.5 2.2	62.4.2 0.4.2.2 0.6.1.2	ಬ್ಲಗಳಳ	<b>a</b> : d d d	44000			0000	ರರ್ಥರ
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		d 354	Ħ	1.70 3.23 2.98 2.55	84.59 9.75 9.75 88.75 9.75 9.75	40000H	0.98 0.98 0.06 0.51	0.55 0.05 0.38 1.67	1.32 1.10 0.79 0.36	0 0 0 0 0 0 0 0 0 0	1.8.4.0.8 3.8.4.0.8 1.0.8	0.00 0.00 0.88 0.88	1.2.2.4.4.	44440 84228	0.25 1.02 1.07 1.07
		Adjusted	н	1.64 3.09 1.88 2.41	82.4.4.8.0 0.82.4.67.28	5. 39 5. 13 5. 87 1. 15	25.05 25.05 25.05 25.05 25.05	00001 88274	0. 13 0. 87 0. 30 0. 16	0.000 8.8.488	5.93 4.74 9.74	0.00 0.00 0.00 0.00 0.00 0.00 0.00 0.0	22.48 2.98.53 171.62	88.89 88.89 88.89	0.10 0.15 0.47 0.70
	ıjab	Adj	П	582 517 096 665 855	164 707 656 237 292	139 0. 24 0. 47 0. 310 0. 408 1.	312 0. 19 0. 160 0. 0 0.	176 273 512 512	233 0. 238 0. 105 0. 57 0.	88 88 88 88	328 996.2. 795.5. 682.4. 056.2.	28 88 88 88 88 88	291 0. 8 518 1. 4 899 2. 8 1058 2. 8 1123 3. 1	1076 1094 1189 566 192	328 288 288 288 288 288
	The Punjab	-	<del> </del>	18632	22 42 15 15 15 15 15 15 15 15 15 15 15 15 15		30 30 30		38588	35548	888218	\$818 <b>\$</b>	<del>284</del> 88	84 118 05 27	24888
	Ţ	d 354.	)H	88525 111111	25248 25150	28.82 17.00.00 1.1.00.00	88288 -000000		82 85 61 61 61 61	23.14.15.0 2.1.1.1.0	08.00 08.00 09.00 09.00	2222 2222 21111	413 13 13 13 13 13 13 13 13 13 13 13 13 1	872.46 11.12.1	25 20 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.
		uste	H	95 1 2 3 1 2 1 3 1 3 1 3 1 3 1 3 1 3 1 3 1	482 177 177 100 100 100 100	82228 -000-1-1	82.488 99999	59 0. 27 1. 38 1. 95 1. 44 0.	54282	15.0 11.1.0 4.3.2.1.0	814888 48166	878858 619868	888818 41446	688888 8.∹9.446	2222 282 2017 2017 2017
		Unadjusted		988 2. 096 3. 920 2. 379 1.	592 4. 552 4. 769 2. 186 0.	40. 800. 1560. 3340. 3771.	293 63 0. 14 0. 8 0. 27 0.	2080.1 4481.3 4881.3 3380.1 1550.4	247 0. 499 1. 782 2 493 1. 181 0.	53 0. 765 2. 077 3. 209 3.	25223 20223 20223 2002	4087.0 4087.0 957.2 937.2	5611. 2950. 5301. 11053.	8512 3470. 5871. 9302.	3551. 2470. 9842. 5094.
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8		78.33	< `│	0.81 1.11 1.12 1.31 1.19	0.011158 0.84598	2555 3555 3555 3555 3555 3555 3555	000		31.85 30.82 31.24 32.18	28282 2883 2883	66.2.16 77.0.88 86.93 841.36	8.8.8.8 8.8.8.8	3.25.27	4444	44486
British Isles	1922		H,	0.65 1.22 1.33 1.43	0,092 0,055 0,055 0,055	3.49 3.49 1.10 0.56	0.45 0.48 0.51 1.48 3.57	8.1.4.4.8 1.1.2.9.8	3, 43 0, 68 1, 53 7, 76	7,59 4,01 3,23 5,16	4,001. 21:22:24	001.08 82883	5.15 1.45 9.57	3225 223 233 233 233 233	7.655.95 7.635.95 7.61
ritis	9850	nadjusted	н	0.63 1.19 1.64 1.33	9.53 9.53 9.53 9.53 9.53	528.83	4.0.5. 4.3.3. 7.1.7.	34.8.8.8.8.8.3.55.0.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.3.55.0.0	21.02.0 21.03.0 20.03.0	6.30 4.20 4.20 7.20 7.10 7.10 7.10	20011 2022 348	6.9.54 2.9.85 2.4.85	3.1.78 3.55 6.81	25.45.2 39.25.23	8888 888 887 88
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Table 5.—Rainfall periodogram, 21/8 to 9 years—(Continued)

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